

2.1 Continuous Parameter Martingales.

(Ω, \mathcal{B}, P) is a probability space and for $t \in [0, T]$, $\mathcal{B}_t \subset \mathcal{B}$ is an increasing family of sub- σ fields, referred to as "filtration". A martingale with respect to $(\Omega, \mathcal{B}_t, P)$ is a family $\xi(t, \omega)$ with the following properties.

- For almost all ω , $\xi(t)$ is a right continuous function of t .
- For each t , $\xi(t, \omega)$ is \mathcal{B}_t measurable. With right continuity it follows that $\xi(\cdot, \omega)$ is "progressively measurable" i.e for each $t > 0$, the function $\xi(s, \omega)$ as a map of $[0, t] \times \Omega \rightarrow R$ is measurable with respect to $\mathcal{B}[0, t] \times \mathcal{B}_t$ where $\mathcal{B}[0, t]$ is the Borel σ -field of $[0, t]$.
- $\xi(t, \omega) \in L_1(P)$ and for $t > s \geq 0$, $E[\xi(t)|\mathcal{B}_s] = \xi(s, \omega)$ a.e.

Remark 2.1. According to a theorem of Doob, a continuous parameter martingale, almost surely, has limits from the left and right at every t . To demand that it be right continuous, i.e. to define the value at t as the limit from the right is a matter of normalization.

Remark 2.2. By restricting the martingale $\xi(t, \omega)$ to a discrete subset $\{nh\}$ we will get a discrete parameter martingale. The usual estimates valid for martingales are valid for them, uniformly in h . We can then let $h \rightarrow 0$ and deduce analogous results for continuous parameter martingales. For example the following theorems are easily established in this manner.

Theorem 2.1. *Let $\xi(t)$ be a continuous parameter martingale on $[0, T]$. Then*

$$P\left[\sup_{t \in [0, T]} |\xi(t)| \geq \ell\right] \leq \frac{1}{\ell} \int_{\{\sup_{t \in [0, T]} |\xi(t)| \geq \ell\}} |\xi(T)| dP \leq \frac{1}{\ell} E[|\xi(T)|]$$

Moreover for $p > 1$,

$$\left\| \sup_{t \in [0, T]} |\xi(t)| \right\|_p \leq \frac{p}{p-1} \|\xi(T)\|_p$$

2.2 Stopping Times.

Given a filtration $\{\mathcal{B}_t\}$ we can define a stopping time relative to the filtration. A function $\tau : \Omega \rightarrow [0, \infty]$ is called a stopping time if for every $t \geq 0$ the set $\{\omega : \tau(\omega) \leq t\}$ is \mathcal{B}_t measurable. Typical examples of stopping times are the first time some thing happens, like the exit time from an open set. Given a stopping time τ there is a natural sub σ -field \mathcal{B}_τ associated with it, defined by

$$A \in \mathcal{B}_\tau \Leftrightarrow A \cap \{\tau \leq t\} \in \mathcal{B}_t \quad \forall t$$

It is easy to check that any stopping time τ is measurable with respect to \mathcal{B}_τ . For any t , $\tau \wedge t$ is a stopping time as well, and $\mathcal{B}_{\tau \wedge t}$ is a new filtration. If $\xi(t)$ is a martingale with respect to $\{\mathcal{B}_t\}$ so is $\xi(\tau \wedge t, \omega)$ with respect to $\mathcal{B}_{\tau \wedge t}$.

Doob's optional stopping theorem for martingales extends to the continuous case.

Theorem 2.2. *If $0 \leq \tau_1 \leq \tau_2 \leq C$ are two bounded stopping times, and $\xi(t)$ is a martingale with respect to $(\Omega, \mathcal{B}_t, P)$ then almost surely*

$$E[\xi(\tau_2)|\mathcal{B}_{\tau_1}] = \xi(\tau_1)$$

This is proved by approximating the stopping times $\tau_i, i = 1, 2$ by $\tau_i^n = \frac{[n\tau_i]+1}{n}$. Then the optional stopping theorem can be applied to the discrete martingale $\xi(\frac{\cdot}{n})$, conclude that $E[\xi(\tau_2^n)|\mathcal{B}_{\tau_1^n}] = \xi(\tau_1^n)$ and let $n \rightarrow \infty$ to obtain our theorem.

2.3 Strong Markov Property.

Brownian motion is a process with independent increments. It is therefore, in particular, a Markov Process. That is to say, given the past history \mathcal{B}_s , [the σ -field generated by $\{x(u) : 0 \leq u \leq s\}$], the conditional distribution of $x(t) = x(s) + [x(t) - x(s)]$ for $t > s$ is the normal distribution with mean $x(s)$ and variance $t - s$. Since this only depends on $x(s)$ the Markov property holds. The strong Markov property extends this from constant times s to stopping times.

We begin with $(\Omega, \mathcal{B}_t, x(t), P)$, where \mathcal{B}_t is an increasing family of sub- σ fields, satisfying

- For each t , $x(t, \omega)$ is \mathcal{B}_t measurable
- $x(t)$ is almost surely a continuous function of t .
- For $t > s >$ almost surely

$$P[x(t) \in A | \mathcal{B}_s] = \int_A \frac{1}{\sqrt{2\pi(t-s)}} e^{-\frac{(y-x(s))^2}{2(t-s)}} dy$$

We will call such an $x(t)$ a Brownian motion adapted to $\{\mathcal{B}_t\}$. Note that \mathcal{B}_t can be larger than $\sigma\{x(s) : 0 \leq s \leq t\}$.

Theorem 2.3. *The strong Markov property holds for Brownian Motion. That is, given any stopping time τ that is almost surely finite,*

$$P[x(t+\tau) \in A | \mathcal{B}_\tau] = \int_A \frac{1}{\sqrt{2\pi t}} e^{-\frac{(y-x(\tau))^2}{2t}} dy$$

Equivalently the process $x(t+\tau) - x(\tau)$ is another Brownian Motion adapted to $\mathcal{B}_{\tau+t}$, and is independent of \mathcal{B}_τ .

Proof. It is enough to show that if $A \in \mathcal{B}_\tau$ and f is a bounded continuous function, then

$$\int_A f(x(t+\tau)) dP = \int_A g(x(\tau)) dP$$

where

$$g(x) = \int f(y) \frac{1}{\sqrt{2\pi t}} e^{-\frac{(y-x)^2}{2t}} dy$$

We will check it for τ that takes only a countable set of values. $\tau = t_j$ for some j . The set $A_j = A \cap \{\tau = t_j\} \in \mathcal{B}_{t_j}$. Therefore from the Markov property

$$\begin{aligned} \int_A f(x(t+\tau)) dP &= \sum_j \int_{A_j} f(x(t+\tau)) dP = \sum_j \int_{A_j} f(x(t+t_j)) dP \\ &= \sum_j \int_{A_j} g(x(t_j)) dP = \int_A g(x(\tau)) dP \end{aligned}$$

If we now approximate τ by $\tau_n = \frac{[n\tau]+1}{n}$ and pass to the limit we are done. Note that here we approximate τ by $\tau_n \geq \tau$, so that $A \in \mathcal{B}_\tau \subset \mathcal{B}_{\tau_n}$. We have also used the fact that g is continuous. \square

Remark 2.3. Any Markov process that has almost surely right continuous paths and $E[f(x(t)|x(s))] = g(s, t, x(s))$ where $g(s, t, x)$ is continuous in x for each fixed $s < t$, has the strong Markov property by the same argument.

2.4 Reflection Principle.

If $x(t)$ is Brownian motion

$$P[\sup_{0 \leq s \leq t} x(s) \geq \ell] = 2P[x(t) \geq \ell]$$

Let τ be the stopping time $\tau = \inf\{s : x(s) \geq \ell\}$. We are interested in $P[\tau \leq t]$. Note that $x(\tau) = \ell$. Therefore

$$\begin{aligned} P[x(t) \geq \ell] &= P[x(t) \geq \ell \ \& \ \tau \leq t] = P[\tau \leq t] \int_\ell^\infty \frac{1}{\sqrt{2\pi(t-\tau)}} e^{-\frac{(y-\ell)^2}{2(t-\tau)}} dy \\ &= \frac{1}{2} P[\tau \leq t] \end{aligned}$$

2.5 Brownian Motion as a Martingale

P is the Wiener measure on (Ω, \mathcal{B}) where $\Omega = C[0, T]$ and \mathcal{B} is the Borel σ -field on Ω . In addition we denote by \mathcal{B}_t the σ -field generated by $x(s)$ for $0 \leq s \leq t$. It is easy to see that $x(t)$ is a martingale with respect to $(\Omega, \mathcal{B}_t, P)$, i.e. for each $t > s$ in $[0, T]$

$$E^P[x(t)|\mathcal{B}_s] = x(s) \quad \text{a.e. } P \quad (2.1)$$

and so is $x(t)^2 - t$. In other words

$$E^P[x(t)^2 - t | \mathcal{F}_s] = x(s)^2 - s \quad \text{a.e. } P \quad (2.2)$$

The proof is rather straight forward. We write $x(t) = x(s) + Z$ where $Z = x(t) - x(s)$ is a random variable independent of the past history \mathcal{B}_s and is distributed as a Gaussian random variable with mean 0 and variance $t - s$. Therefore $E^P[Z|\mathcal{B}_s] = 0$ and $E^P[Z^2|\mathcal{B}_s] = t - s$ a.e P . Conversely,

Theorem 2.4. Lévy's theorem. *If P is a measure on $(C[0, T], \mathcal{B})$ such that $P[x(0) = 0] = 1$ and the functions $x(t)$ and $x^2(t) - t$ are martingales with respect to $(C[0, T], \mathcal{B}_t, P)$ then P is the Wiener measure.*

Proof. The proof is based on the observation that a Gaussian distribution is determined by two moments. But that the distribution is Gaussian is a consequence of the fact that the paths are almost surely continuous and not part of our assumptions. The actual proof is carried out by establishing that for each real number λ

$$X_\lambda(t) = \exp \left[\lambda x(t) - \frac{\lambda^2}{2} t \right] \quad (2.3)$$

is a martingale with respect to $(C[0, T], \mathcal{B}_t, P)$. Once this is established it is elementary to compute

$$E^P \left[\exp \left[\lambda (x(t) - x(s)) \right] | \mathcal{B}_s \right] = \exp \left[\frac{\lambda^2}{2} (t - s) \right]$$

which shows that we have a Gaussian Process with independent increments with two matching moments. The proof of (2.3) is more or less the same as proving the central limit theorem. In order to prove that $X_\lambda(t)$ is a martingale, we can assume with out loss of generality that $s = 0$ and show that

$$E^P \left[\exp \left[\lambda x(t) - \frac{\lambda^2}{2} t \right] \right] = 1 \quad (2.4)$$

To this end let us define successively $\tau_{0,\epsilon} = 0$,

$$\tau_{k+1,\epsilon} = \min \left[\inf \{ s : s \geq \tau_{k,\epsilon}, |x(s) - x(\tau_{k,\epsilon})| \geq \epsilon \}, t, \tau_{k,\epsilon} + \epsilon \right]$$

Then each $\tau_{k,\epsilon}$ is a stopping time and eventually $\tau_{k,\epsilon} = t$ by continuity of paths. The continuity of paths also guarantees that $|x(\tau_{k+1,\epsilon}) - x(\tau_{k,\epsilon})| \leq \epsilon$. We write

$$x(t) = \sum_{k \geq 0} [x(\tau_{k+1,\epsilon}) - x(\tau_{k,\epsilon})]$$

and

$$t = \sum_{k \geq 0} [\tau_{k+1,\epsilon} - \tau_{k,\epsilon}]$$

To establish (2.4) we calculate the quantity on the left hand side as

$$\lim_{n \rightarrow \infty} E^P \left[\exp \left[\sum_{0 \leq k \leq n} \left[\lambda [x(\tau_{k+1,\epsilon}) - x(\tau_{k,\epsilon})] - \frac{\lambda^2}{2} [\tau_{k+1,\epsilon} - \tau_{k,\epsilon}] \right] \right] \right]$$

and show that it is equal to 1. Let us consider the σ -field $\mathcal{F}_k = \mathcal{B}_{\tau_{k,\epsilon}}$ and the quantity

$$q_k(\omega) = E^P \left[\exp \left[\lambda [x(\tau_{k+1,\epsilon}) - x(\tau_{k,\epsilon})] - \left[\frac{\lambda^2}{2} + C(\lambda)\epsilon \right] [\tau_{k+1,\epsilon} - \tau_{k,\epsilon}] \right] \middle| \mathcal{F}_k \right]$$

with the choice of the constant $C(\lambda)$ to be chosen later. Clearly, if we use Taylor expansion and the fact that $x(t)$ as well as $x(t)^2 - t$ are martingales

$$\begin{aligned} q_k(\omega) &\leq E^P \left[1 + c(\lambda) \left[|x(\tau_{k+1,\epsilon}) - x(\tau_{k,\epsilon})|^3 + |\tau_{k+1,\epsilon} - \tau_{k,\epsilon}|^2 \right] - C(\lambda)\epsilon [\tau_{k+1,\epsilon} - \tau_{k,\epsilon}] \middle| \mathcal{F}_k \right] \\ &\leq E^P \left[1 + c(\lambda)\epsilon \left[|x(\tau_{k+1,\epsilon}) - x(\tau_{k,\epsilon})|^2 + |\tau_{k+1,\epsilon} - \tau_{k,\epsilon}| \right] - C(\lambda)\epsilon [\tau_{k+1,\epsilon} - \tau_{k,\epsilon}] \middle| \mathcal{F}_k \right] \\ &\leq 1 \end{aligned}$$

for some suitably chosen constant $C(\lambda)$ depending on λ . By Fatou's lemma

$$E^P \left[\exp \left[\lambda x(t) - \left[\frac{\lambda^2}{2} + C(\lambda)\epsilon \right] t \right] \right] \leq 1$$

Since $\epsilon > 0$ is arbitrary we prove one half of (2.4). A similar estimate will yield

$$E^P \left[\exp \left[\lambda x(t) - \left[\frac{\lambda^2}{2} - C(\lambda)\epsilon \right] t \right] \right] \geq 1$$

which can be used to prove the other half provided we show the uniform integrability of $\{\exp[\lambda x(\tau_n)]\}$. This follows from the upper bound established above. This completes the proof of the theorem. \square

Remark 2.4. Theorem 2.4 fails for the process $x(t) = N(t) - t$ where $N(t)$ is the standard Poisson Process with rate 1.

Remark 2.5. One can use the Martingale inequality in order to estimate the probability $P\{\sup_{0 \leq s \leq t} |x(s)| \geq \ell\}$. For $\lambda > 0$, by Doob's inequality

$$P \left[\sup_{0 \leq s \leq t} \exp \left[\lambda x(s) - \frac{\lambda^2}{2} s \right] \geq A \right] \leq \frac{1}{A}$$

and

$$\begin{aligned} P \left[\sup_{0 \leq s \leq t} x(s) \geq \ell \right] &\leq P \left[\sup_{0 \leq s \leq t} \left[x(s) - \frac{\lambda s}{2} \right] \geq \ell - \frac{\lambda t}{2} \right] \\ &= P \left[\sup_{0 \leq s \leq t} \left[\lambda x(s) - \frac{\lambda^2 s}{2} \right] \geq \lambda \ell - \lambda^2 t / 2 \right] \\ &\leq \exp \left[-\lambda \ell + \frac{\lambda^2 t}{2} \right] \end{aligned}$$

Optimizing over $\lambda > 0$, we obtain

$$P \left[\sup_{0 \leq s \leq t} x(s) \geq \ell \right] \leq \exp \left[-\frac{\ell^2}{2t} \right]$$

and by symmetry

$$P\left[\sup_{0 \leq s \leq t} |x(s)| \geq \ell\right] \leq 2 \exp\left[-\frac{\ell^2}{2t}\right]$$

The estimate is not too bad because by reflection principle

$$P\left[\sup_{0 \leq s \leq t} x(s) \geq \ell\right] = 2P[x(t) \geq \ell] = \sqrt{\frac{2}{\pi t}} \int_{\ell}^{\infty} \exp\left[-\frac{x^2}{2t}\right] dx$$

Exercise 2.1. One can use the estimate above to prove the result of Paul Lévy

$$P\left[\limsup_{\delta \rightarrow 0} \frac{\sup_{\substack{0 \leq s, t \leq 1 \\ |s-t| \leq \delta}} |x(s) - x(t)|}{\sqrt{\delta \log \frac{1}{\delta}}} = \sqrt{2}\right] = 1$$

We had an exercise in the previous section that established the lower bound. Let us concentrate on the upper bound. If we define

$$\Delta_{\delta}(\omega) = \sup_{\substack{0 \leq s, t \leq 1 \\ |s-t| \leq \delta}} |x(s) - x(t)|$$

first check that it is sufficient to prove that for any $\rho < 1$, and $a > \sqrt{2}$

$$\sum_n P[\Delta_{\rho^n}(\omega) \geq a \sqrt{n\rho^n \log \frac{1}{\rho}}] < \infty \quad (2.5)$$

To estimate $\Delta_{\rho^n}(\omega)$ it is sufficient to estimate $\sup_{t \in I_j} |x(t) - x(t_j)|$ for $k_{\epsilon} \rho^{-n}$ overlapping intervals $\{I_j\}$ of the form $[t_j, t_j + (1 + \epsilon)\rho^n]$ with length $(1 + \epsilon)\rho^n$. For each $\epsilon > 0$, $k_{\epsilon} = \epsilon^{-1}$ is a constant such that any interval $[s, t]$ of length no larger than ρ^n is completely contained in some I_j with $t_j \leq s \leq t_j + \epsilon\rho^n$. Then

$$\Delta_{\rho^n}(\omega) \leq \sup_j \left[\sup_{t \in I_j} |x(t) - x(t_j)| + \sup_{t_j \leq s \leq t_j + \epsilon\rho^n} |x(s) - x(t_j)| \right]$$

Therefore, for any $a = a_1 + a_2$,

$$\begin{aligned} & P\left[\Delta_{\rho^n}(\omega) \geq a \sqrt{n\rho^n \log \frac{1}{\rho}}\right] \\ & \leq P\left[\sup_j \sup_{t \in I_j} |x(t) - x(t_j)| \geq a_1 \sqrt{n\rho^n \log \frac{1}{\rho}}\right] \\ & \quad + P\left[\sup_j \sup_{t_j \leq s \leq t_j + \epsilon\rho^n} |x(s) - x(t_j)| \geq a_2 \sqrt{n\rho^n \log \frac{1}{\rho}}\right] \\ & \leq 2k_{\epsilon} \rho^{-n} \left[\exp\left[-\frac{a_1^2 n\rho^n \log \frac{1}{\rho}}{2(1 + \epsilon)\rho^n}\right] + \exp\left[-\frac{a_2^2 n\rho^n \log \frac{1}{\rho}}{2\epsilon\rho^n}\right] \right] \end{aligned}$$

Since $a > \sqrt{2}$, we can pick $a_1 > \sqrt{2}$ and $a_2 > 0$. For $\epsilon > 0$ sufficiently small (2.5) is easily verified.